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Networks from correlation matrices: An alternative to thresholding

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**[Moving to State University of New York at Buffalo,
Department of Mathematics, mid August]**

Threshold Networks
University of Nottingham
22 July 2019

A typical work flow

e.g. schizophrenia

(Lynall, Bassett, Kerwin, McKenna, Kitzbichler, Muller & Bullmore, JNS 2010)

Correlation matrix

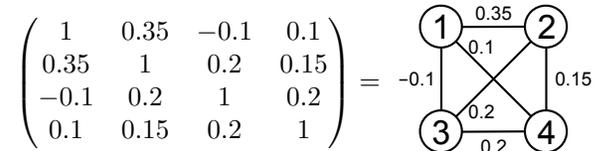
- A major form of multivariate data
- Connection between i and j = strength of correlation between i and j
- Pearson (or other) correlation
- Measurements: time series, feature vector, questionnaires across human/animal subjects etc.
 - Financial time series, neuroimaging, genetic data, psychology experiments, climate networks etc.
- Conventional approaches
 - Principal component analysis
 - Factor analysis
 - Random matrix theory
 - Markowitz's optimal portfolio

Correlation networks

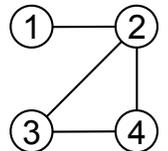
- To benefit from network analysis tools
- Two conventional approaches

1. Thresholding

2. Weighted networks



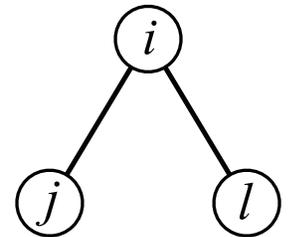
correlation matrix



thresholded network

- Brain networks, financial networks, ...
- Problems:

- False positives (pseudo correlation)
- Loss of information by thresholding
- Arbitrariness of the threshold choice
- Negative correlation in the case of weighted networks



Grand goal

- Analyse correlation matrices as they are, and also exploit various network concepts/algorithms
- Without thresholding or discarding negatively correlated node pairs

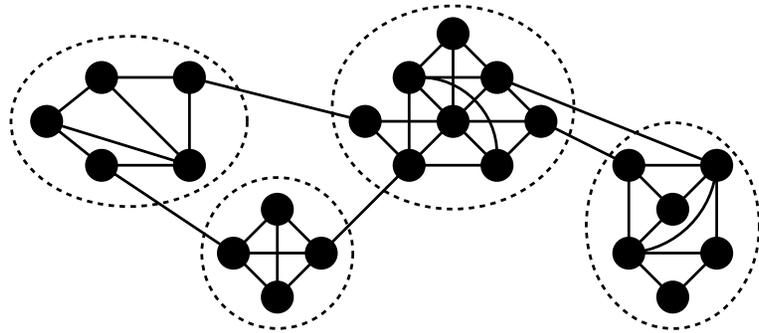
PHYSICAL REVIEW X **5**, 021006 (2015)

Community Detection for Correlation Matrices

Mel MacMahon¹ and Diego Garlaschelli^{1,2}

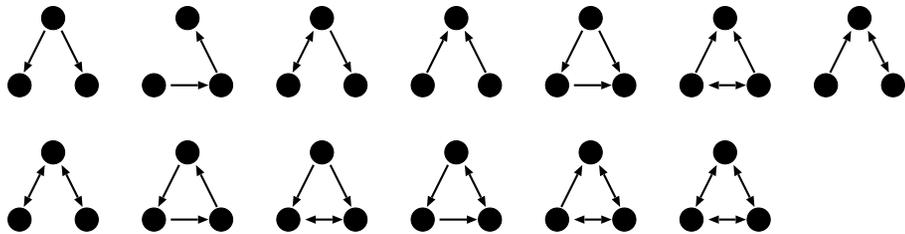
- Null model = identity matrix, or random matrices based on random matrix theory

- However, we usually say that a network has a property α relative to the “configuration model” (degree-preserving random graph)

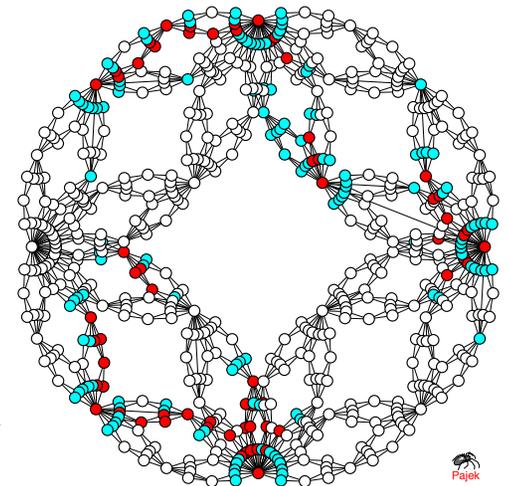


community structure

“rich clubs” Colizza et al. (2006)



network motifs



fractality

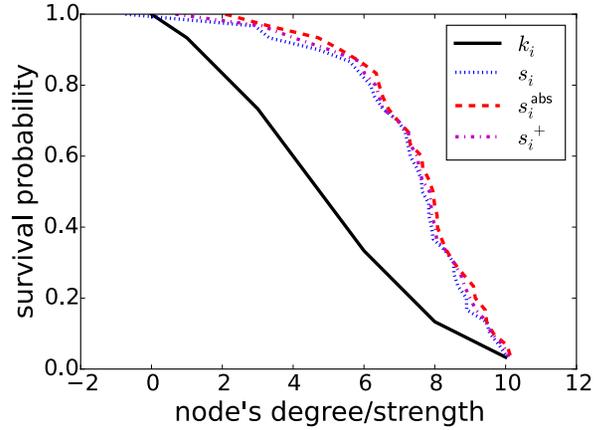
Aim

- Propose a configuration model for correlation matrices
- Strategy: maximum entropy
- Then, use it for community detection, measurement of clustering coefficients, network filtering etc., hopefully to enrich analyses of correlation matrix data

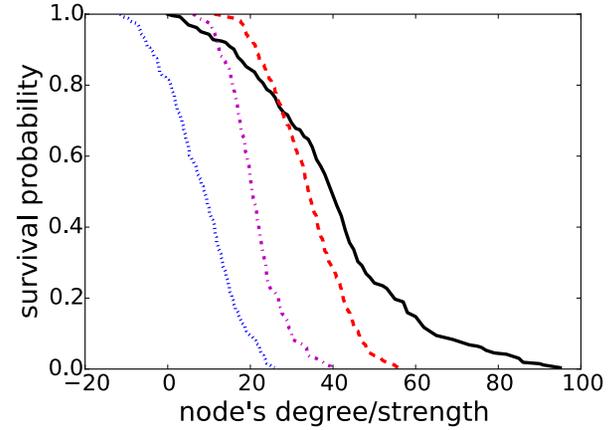
Methods in short

- Distribution of correlation/covariance matrices
- Boltzmann distribution (exponential family)
 - cf. exponential random graph models
- Constraints
 - Weighted (and signed) degree of each node is conserved.

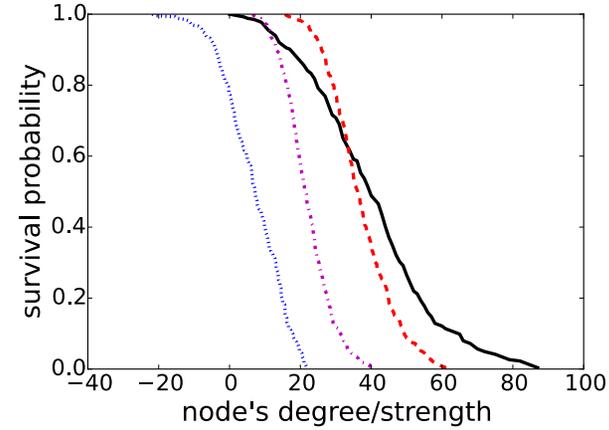
(a) motivation



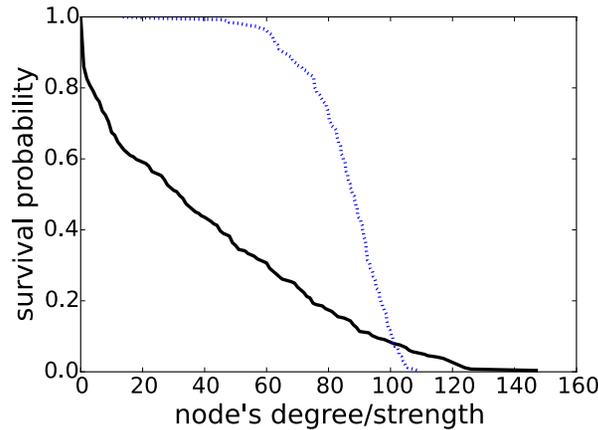
(b) fMRI1



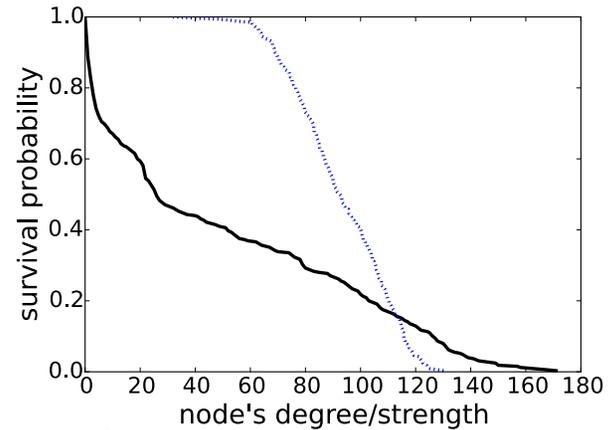
(c) fMRI2



(d) Japanese stocks



(e) US stocks



$$s_i = \sum_{j=1; j \neq i}^N \text{Cor}_{ij}$$

$$s_i^{\text{abs}} = \sum_{j=1; j \neq i}^N |\text{Cor}_{ij}|$$

$$s_i^+ = \sum_{j=1; j \neq i; \text{Cor}_{ij} > 0}^N \text{Cor}_{ij}$$

$$s_i \approx s_i^{\text{abs}} \approx s_i^+$$

because $\text{Cor}_{ij}^{\text{org}} \geq 0$ for all but three pairs of nodes

✓ Not long-tailed but moderately heterogeneous

- Work on covariance (as opposed to correlation) matrices basically, as that is mathematically easier.

$$\text{Cor}_{ij} = \frac{\text{Cov}_{ij}}{\sqrt{\text{Cov}_{ii}\text{Cov}_{jj}}}$$

$$\text{Cov}^{\text{con}} = (\text{Cov}_{ij}^{\text{con}}) = \frac{1}{L} X X^{\top}$$

Equivalently

$$\text{Cov}_{ij}^{\text{con}} = \frac{1}{L} \sum_{\ell=1}^L x_{i\ell} x_{j\ell}$$

And consider the distribution of X

Rationales:

1. Realistic
2. Mathematically feasible

Maxent preserving node strength

$$H(X) \equiv - \int p(X) \ln p(X) dX + \sum_{i=1}^N \alpha_i \left[\int \text{Cov}_{ii}^{\text{con}} p(X) dX - \text{Cov}_{ii}^{\text{org}} \right] \\ + \sum_{i=1}^N \beta_i \left[\int \sum_{j=1; j \neq i}^N \text{Cov}_{ij}^{\text{con}} p(X) dX - \sum_{j=1; j \neq i}^N \text{Cov}_{ij}^{\text{org}} \right],$$

$$p(X) \propto \exp \left[\frac{1}{L} \sum_{\ell=1}^L \left(\sum_{i=1}^N \alpha_i x_{i\ell}^2 + \sum_{i=1}^N \beta_i \sum_{j=1; j \neq i}^N x_{i\ell} x_{j\ell} \right) \right]$$

$$= \prod_{\ell=1}^L \exp \left[\frac{1}{L} \left(\sum_{i=1}^N \alpha_i x_{i\ell}^2 + \sum_{i=1}^N \beta_i \sum_{j=1; j \neq i}^N x_{i\ell} x_{j\ell} \right) \right]$$

$$= \prod_{\ell=1}^L \exp \left[-\frac{1}{2} \mathbf{x}_\ell^\top \Sigma^{-1} \mathbf{x}_\ell \right],$$

Reminder: $\text{Cov}^{\text{con}} = (\text{Cov}_{ij}^{\text{con}}) = \frac{1}{L} X X^\top$

$$\text{Cov}_{ij}^{\text{con}} = \frac{1}{L} \sum_{\ell=1}^L x_{i\ell} x_{j\ell}$$

$$p(X) \propto \prod_{\ell=1}^L \exp \left[-\frac{1}{2} \mathbf{x}_\ell^\top \Sigma^{-1} \mathbf{x}_\ell \right]$$

where $\mathbf{x}_\ell = (x_{1\ell}, \dots, x_{N,\ell})^\top$ and

$$\Sigma^{-1} = -\frac{1}{L} \begin{pmatrix} 2\alpha_1 & \beta_1 + \beta_2 & \beta_1 + \beta_3 & \cdots & \beta_1 + \beta_{N-1} & \beta_1 + \beta_N \\ \beta_2 + \beta_1 & 2\alpha_2 & \beta_2 + \beta_3 & \cdots & \beta_2 + \beta_{N-1} & \beta_2 + \beta_N \\ & & \vdots & & & \\ \beta_N + \beta_1 & \beta_N + \beta_2 & \cdots & \cdots & \beta_N + \beta_{N-1} & 2\alpha_N \end{pmatrix}$$

So multivariate normal distribution:

$$p(X) = \prod_{\ell=1}^L \frac{1}{\sqrt{(2\pi)^N |\Sigma|}} \exp \left[-\frac{1}{2} \mathbf{x}_\ell^\top \Sigma^{-1} \mathbf{x}_\ell \right]$$

To infer parameters

- Gradient descent method

$$\alpha_i^{\text{new}} = \alpha_i^{\text{old}} + \epsilon (\Sigma_{ii} - \text{Cov}_{ii}^{\text{org}}),$$

$$\beta_i^{\text{new}} = \beta_i^{\text{old}} + 2\epsilon \left(\sum_{j=1}^N \Sigma_{ij} - \sum_{j=1}^N \text{Cov}_{ij}^{\text{org}} \right),$$

- Convex optimisation
 - Faster and more accurate
 - Documentation and code on Github (not in the first PRE paper but in the second paper on arXiv)



Sadamori Kojaku

Numerical simulations: Aim

- Strength (i.e. weighted degree) of each node really preserved?
- Speed
- Comparison with other null models for correlation matrices (e.g. MacMahon & Garlaschelli, PRX 2015)
- Deployment to algorithms or statistical tests that need a null model

Data

- Psychological questionnaires; $N=30$ questions, $L=686$ participants
- Functional MRI from healthy participants; $N=264$ locations, $L=4760$ time points
- Stocks, Japan
 - $N=264$ stocks belonging to the first section of the Tokyo Stock Exchange provided by Nikkei NEEDS
 - from 12 March 1996 to 29 February 2016
 - $L=4904$ trading days (daily observations)
- Stocks, US
 - $N=325$ stocks from Standard & Poor's 500 index using MATHEMATICA FinancialData package
 - from 3 January 1996 to 24 February 2017
 - $L=5324$ trading days

Yukie Sano



H-Q-S null model

- Hirschberger-Qu-Steuer, Eur J Oper Res, 177, 1610 (2007)
- For covariance matrices
- Preserves
 - mean of the on-diagonal entries
 - mean of the off-diagonal entries
 - variance of the off-diagonal entries
- Corresponds to Erdős-Rényi random graph

Null models by MacMahon & Garlaschelli, PRX 2015

Marcenko-Pastur (also called Senguta-Mitra) distribution:

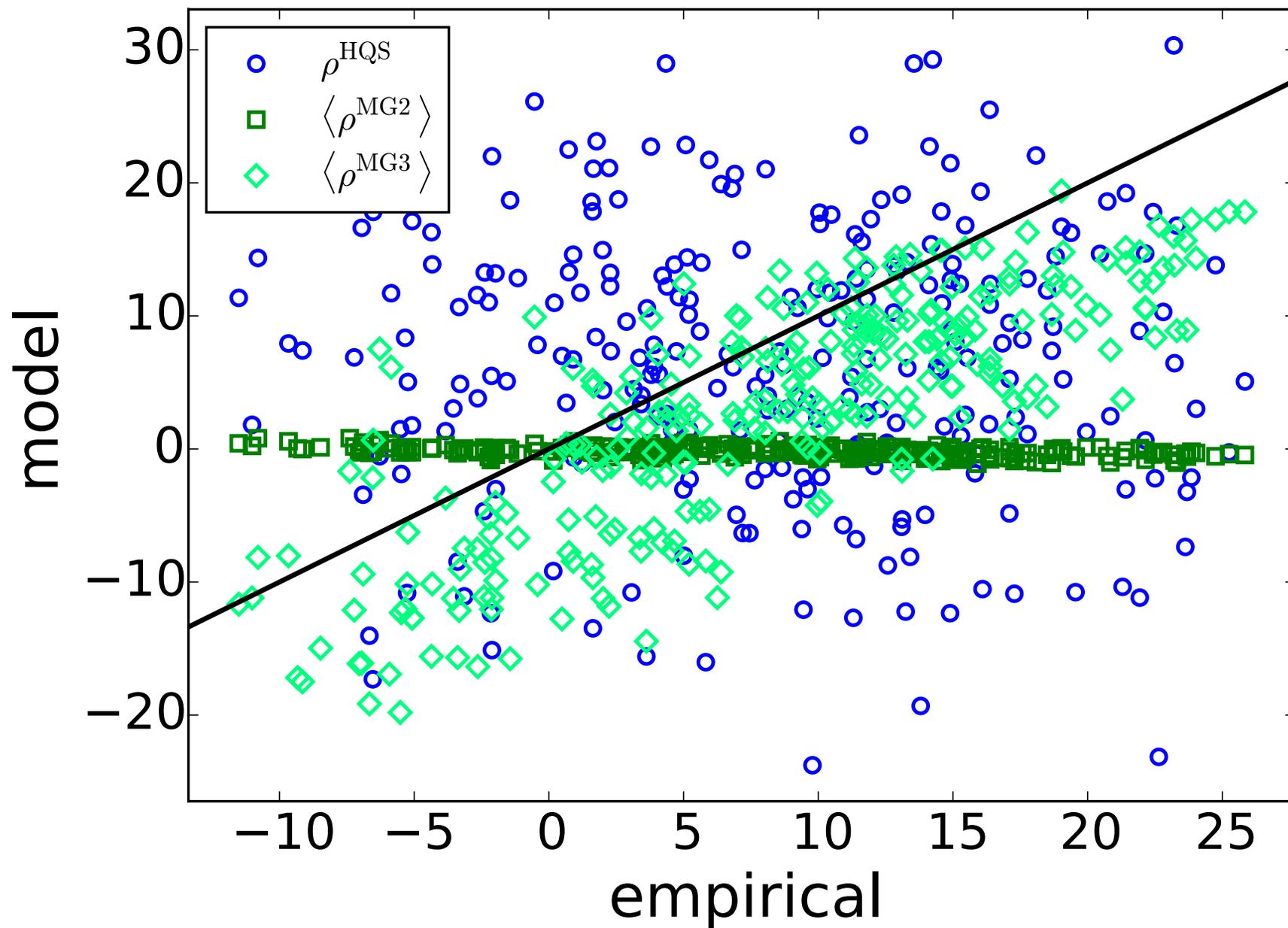
$$\rho(\lambda) = \begin{cases} \frac{L}{N} \frac{\sqrt{(\lambda_+ - \lambda)(\lambda - \lambda_-)}}{2\pi\lambda} & (\lambda_- \leq \lambda \leq \lambda_+), \\ 0 & (\text{otherwise}), \end{cases}$$

$$\text{where } \lambda_{\pm} = \left(1 \pm \sqrt{N/L}\right)^2$$

- Model MG2: Preserves modes whose eigenvalues $\lambda \in [\lambda_-, \lambda_+]$
- Model MG3: Preserves modes with $\lambda \in [\lambda_-, \lambda_+]$ & λ_{\max}
- Not generating (random) correlation *matrices*

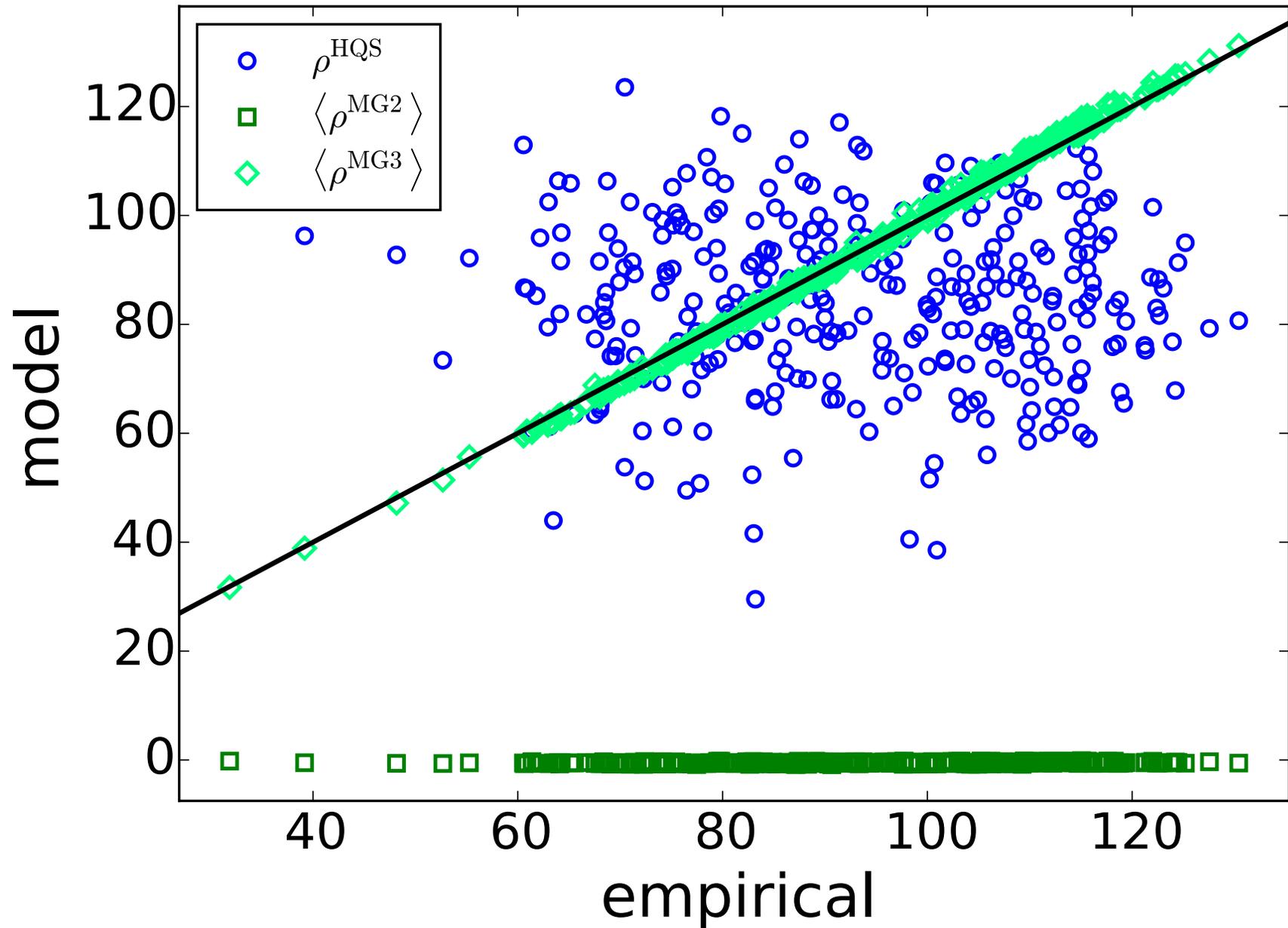
Node strength

fMRI (participant 1)



Node strength

US stocks



Clustering coefficient

Onnela et al. (Phys Rev E 2005)
for weighted networks

Masuda et al. (Frontiers in
Neuroinformatics 2018) for
correlation matrices

$C^{\text{wei},O}$

$C^{\text{cor},M}$

Null model	C	Z	P	C	Z	P
Motivation	0.284 (empirical)			0.031 (empirical)		
Configuration	0.503 ± 0.024	-9.23	$<10^{-3}$	0.022 ± 0.001	6.20	$<10^{-3}$
H-Q-S	0.335 ± 0.034	-1.51	0.130	0.029 ± 0.003	0.59	0.557
White-noise	0.111 ± 0.015	11.26	$<10^{-3}$	0.001 ± 0.000	869.28	$<10^{-3}$
fMRI1	0.096 (empirical)			0.013 (empirical)		
Configuration	0.138 ± 0.004	-10.11	$<10^{-3}$	0.003 ± 0.000	85.90	$<10^{-3}$
H-Q-S	0.127 ± 0.005	-6.29	$<10^{-3}$	0.023 ± 0.000	-59.50	$<10^{-3}$
White-noise	0.078 ± 0.005	3.66	$<10^{-3}$	0.000 ± 0.000	23346.41	$<10^{-3}$
Japan	0.413 (empirical)			0.027 (empirical)		
Configuration	0.613 ± 0.008	-25.38	$<10^{-3}$	0.026 ± 0.001	1.32	0.188
H-Q-S	0.425 ± 0.014	-0.89	0.376	0.024 ± 0.001	3.77	$<10^{-3}$
White-noise	0.077 ± 0.005	62.77	$<10^{-3}$	0.000 ± 0.000	48473.21	$<10^{-3}$
US	0.328 (empirical)			0.024 (empirical)		
Configuration	0.508 ± 0.007	-25.96	$<10^{-3}$	0.023 ± 0.000	2.80	0.005
H-Q-S	0.362 ± 0.012	-2.79	0.005	0.022 ± 0.001	4.82	$<10^{-3}$
White-noise	0.075 ± 0.005	51.94	$<10^{-3}$	0.000 ± 0.000	59738.94	$<10^{-3}$

Community detection

Max.

$$Q = \frac{1}{C_{\text{norm}}} \sum_{i,j=1}^N (\rho_{ij} - \langle \rho_{ij} \rangle) \delta(g_i, g_j)$$

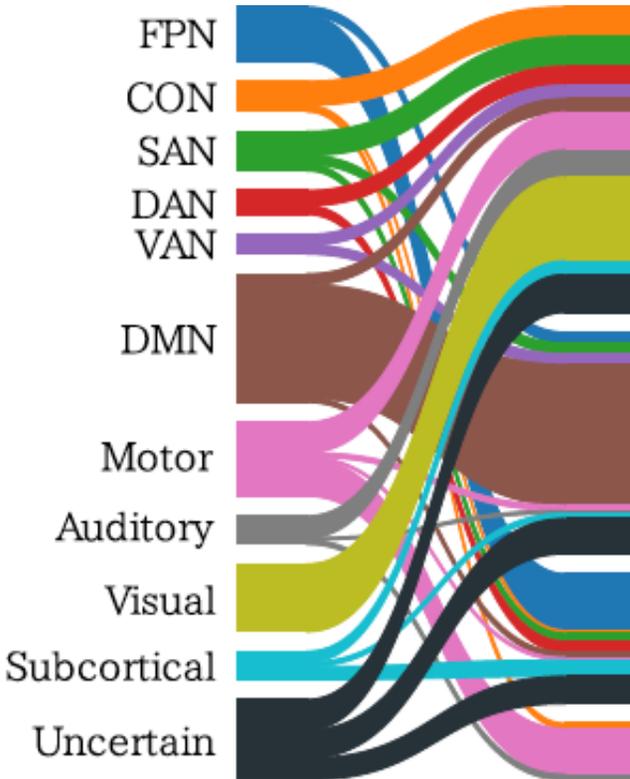
When the configuration model is fed as input,

Null model	Q		Z	P
	Original	Random		
Motivation				
$\langle \rho^{\text{HQS}} \rangle$	0.15	0.09 ± 0.02	2.72	0.007
$\langle \rho^{\text{MG1}} \rangle$	0.88	0.88 ± 0.01	0.85	0.204
$\langle \rho^{\text{MG2}} \rangle$	0.98	0.97 ± 0.00	3.18	$<10^{-3}$
$\langle \rho^{\text{MG3}} \rangle$	0.12	0.22 ± 0.01	-7.00	1.000
fMRI1				
$\langle \rho^{\text{HQS}} \rangle$	0.61	0.17 ± 0.01	35.37	$<10^{-3}$
$\langle \rho^{\text{MG1}} \rangle$	1.10	0.89 ± 0.00	54.62	$<10^{-3}$
$\langle \rho^{\text{MG2}} \rangle$	0.11	1.02 ± 0.00	-966.06	1.000
$\langle \rho^{\text{MG3}} \rangle$	0.11	0.50 ± 0.01	-36.92	1.000
Japan				
$\langle \rho^{\text{HQS}} \rangle$	0.08	0.04 ± 0.01	7.76	$<10^{-3}$
$\langle \rho^{\text{MG1}} \rangle$	0.99	0.99 ± 0.00	-0.19	0.582
$\langle \rho^{\text{MG2}} \rangle$	0.01	1.00 ± 0.00	-25328.61	1.000
$\langle \rho^{\text{MG3}} \rangle$	0.01	0.09 ± 0.00	-17.07	1.000
US				
$\langle \rho^{\text{HQS}} \rangle$	0.10	0.05 ± 0.01	9.29	$<10^{-3}$
$\langle \rho^{\text{MG1}} \rangle$	0.99	0.99 ± 0.00	-0.00	0.482
$\langle \rho^{\text{MG2}} \rangle$	0.01	1.00 ± 0.00	-27404.80	1.000
$\langle \rho^{\text{MG3}} \rangle$	0.01	0.11 ± 0.00	-19.89	1.000

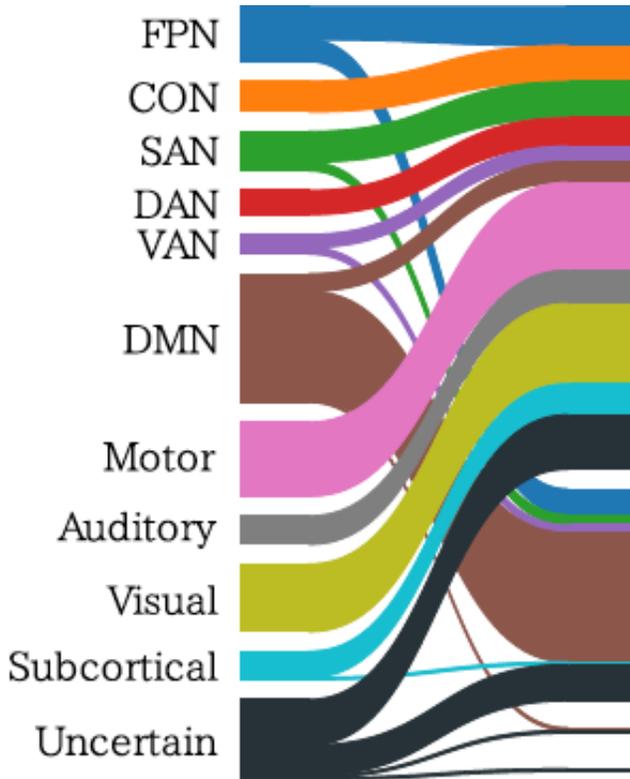
Null model	Q		Z	P
	Original	Random		
Motivation				
$\langle \rho^{\text{con}} \rangle$	0.20	0.03 ± 0.02	10.09	$<10^{-3}$
$\langle \rho^{\text{MG2}} \rangle$	0.99	0.98 ± 0.00	1.58	0.044
$\langle \rho^{\text{MG3}} \rangle$	0.22	0.30 ± 0.02	-3.38	1.000
fMRI1				
$\langle \rho^{\text{con}} \rangle$	0.95	0.05 ± 0.01	83.31	$<10^{-3}$
$\langle \rho^{\text{MG2}} \rangle$	1.82	1.84 ± 0.23	-0.10	0.493
$\langle \rho^{\text{MG3}} \rangle$	1.04	1.71 ± 0.22	-3.01	1.000
Japan				
$\langle \rho^{\text{con}} \rangle$	0.03	0.01 ± 0.01	3.95	0.002
$\langle \rho^{\text{MG2}} \rangle$	1.00	1.00 ± 0.00	4.69	$<10^{-3}$
$\langle \rho^{\text{MG3}} \rangle$	0.04	0.08 ± 0.01	-6.75	1.000
US				
$\langle \rho^{\text{con}} \rangle$	0.05	0.01 ± 0.01	5.13	$<10^{-3}$
$\langle \rho^{\text{MG2}} \rangle$	1.00	1.00 ± 0.00	4.10	$<10^{-3}$
$\langle \rho^{\text{MG3}} \rangle$	0.05	0.11 ± 0.01	-7.91	1.000

functional MRI (participant 1)

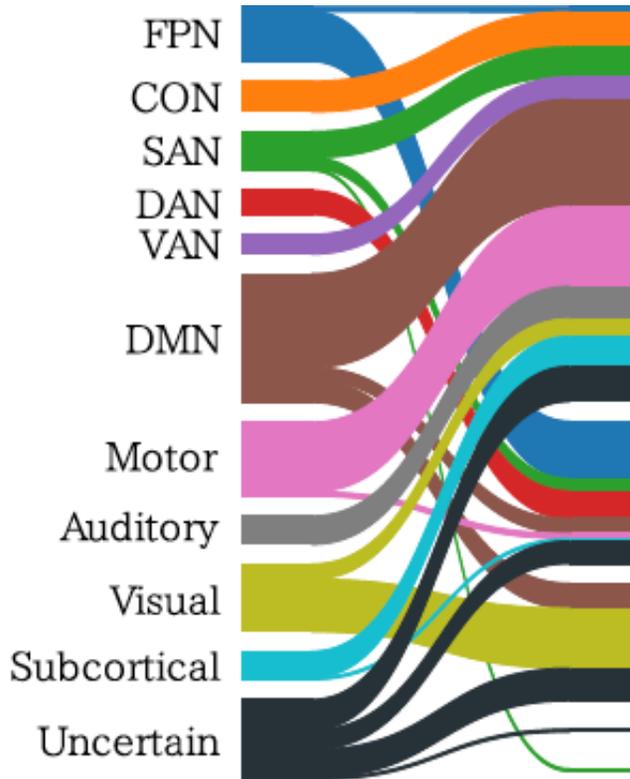
our null model



MG2



MG3



Comparison with biologically determined communities

Null model	Data	p^{emp}	p^{rand}	$\frac{p^{\text{emp}}}{p^{\text{rand}}}$	$\frac{p^{\text{emp}}}{p^{\text{rand}}}$
All nodes					
$\langle \rho^{\text{con}} \rangle$	fMRI1	0.623	0.316	0.306	1.968
	fMRI2	0.673	0.336	0.337	2.004
$\langle \rho^{\text{MG2}} \rangle$	fMRI1	0.725	0.555	0.170	1.307
	fMRI2	0.726	0.520	0.206	1.397
$\langle \rho^{\text{MG3}} \rangle$	fMRI1	0.605	0.434	0.171	1.393
	fMRI2	0.529	0.362	0.167	1.460

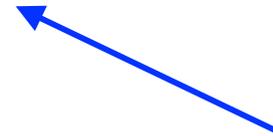
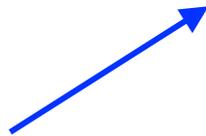
Application to network construction

- We may not be interested in (i, j) if the $\text{corr}(X_i, X_j)$ value is “unsurprising”.
- Statistical tests on the individual edges (node pairs) face a multiple comparison problem. And $\text{corr}(X_1, X_2)$ and $\text{corr}(X_1, X_3)$ are generally correlated with each other.
- Filtering correlation matrices
 - max likelihood + regularization (to estimate sparse networks)

Our contributions

- Main strategies in the field:
 - Estimate sparse precision matrices (“graphical lasso”)
 - Estimate sparse covariance matrices ... seems less popular
- Alternative to “thresholding networks”
- Lacking null model thinking
- Sparse network reconstruction from COrrrelational data with LAzzo, “Scola”
 - Correlation matrix/network
 - Use null models

$$\mathbf{C} = \mathbf{C}^{\text{null}} + \mathbf{W}$$



Estimated corr matrix
(which may not be
equal to the sample
corr matrix)

Null model
calculated from data

Sparse matrix
estimated by lasso

$\mathbf{X} = (x_{\ell i})$ where $\ell = 1, \dots, L; i = 1, \dots, N;$
 N : number of nodes; L : number of samples (e.g. time points in a time series)

ℓ -th sample: $\mathbf{x}_\ell = [x_{\ell 1}, x_{\ell 2}, \dots, x_{\ell N}]$

Sample corr matrix: $\mathbf{C}^{\text{sample}} \equiv \mathbf{X}^\top \mathbf{X} / L$

Given the data and the calculated null model, find \mathbf{W} maximizing

$$P(\mathbf{X} | \mathbf{C}) \equiv \prod_{\ell=1}^L \frac{1}{(2\pi)^{N/2} \det(\mathbf{C})^{1/2}} \exp\left(-\frac{1}{2} \mathbf{x}_\ell \mathbf{C}^{-1} \mathbf{x}_\ell^\top\right)$$

Reminder: $\mathbf{C} = \mathbf{C}^{\text{null}} + \mathbf{W}$

$$P(\mathbf{X}|\mathbf{C}) \equiv \prod_{\ell=1}^L \frac{1}{(2\pi)^{N/2} \det(\mathbf{C})^{1/2}} \exp\left(-\frac{1}{2} \mathbf{x}_\ell \mathbf{C}^{-1} \mathbf{x}_\ell^\top\right)$$

Using $\sum_{\ell=1}^L \mathbf{x}_\ell \mathbf{C}^{-1} \mathbf{x}_\ell^\top = \text{tr}(\mathbf{X}^\top \mathbf{X} \mathbf{C}^{-1})$ we get

$$\ln P(\mathbf{X}|\mathbf{C}) = -\frac{L}{2} \ln \det(\mathbf{C}) - \frac{L}{2} \text{tr}(\mathbf{C}^{\text{sample}} \mathbf{C}^{-1}) - \frac{NL}{2} \ln(2\pi)$$

Using $\mathbf{C} = \mathbf{C}^{\text{null}} + \mathbf{W}$ we get

$$\mathcal{L}(\mathbf{W}) \equiv -\frac{L}{2} \ln \det(\mathbf{C}^{\text{null}} + \mathbf{W}) - \frac{L}{2} \text{tr}\left[\mathbf{C}^{\text{sample}} (\mathbf{C}^{\text{null}} + \mathbf{W})^{-1}\right] - \frac{NL}{2} \ln(2\pi)$$

Though (in fact) $\mathbf{W}^{\text{MLE}} \equiv \mathbf{C}^{\text{sample}} - \mathbf{C}^{\text{null}}$

\mathbf{W}^{MLE} overfits the data, resulting in many (spurious?) edges.

So, use lasso to maximize a penalized (log) likelihood:

$$\hat{\mathcal{L}}(\mathbf{W}|\boldsymbol{\lambda}) \equiv \mathcal{L}(\mathbf{W}) - \frac{L}{2} \sum_{i=1}^N \sum_{j=1}^{i-1} \lambda_{ij} |W_{ij}|;$$

Not concave w.r.t. \mathbf{W} , so use an extension of a previous numerical optimization algorithm.

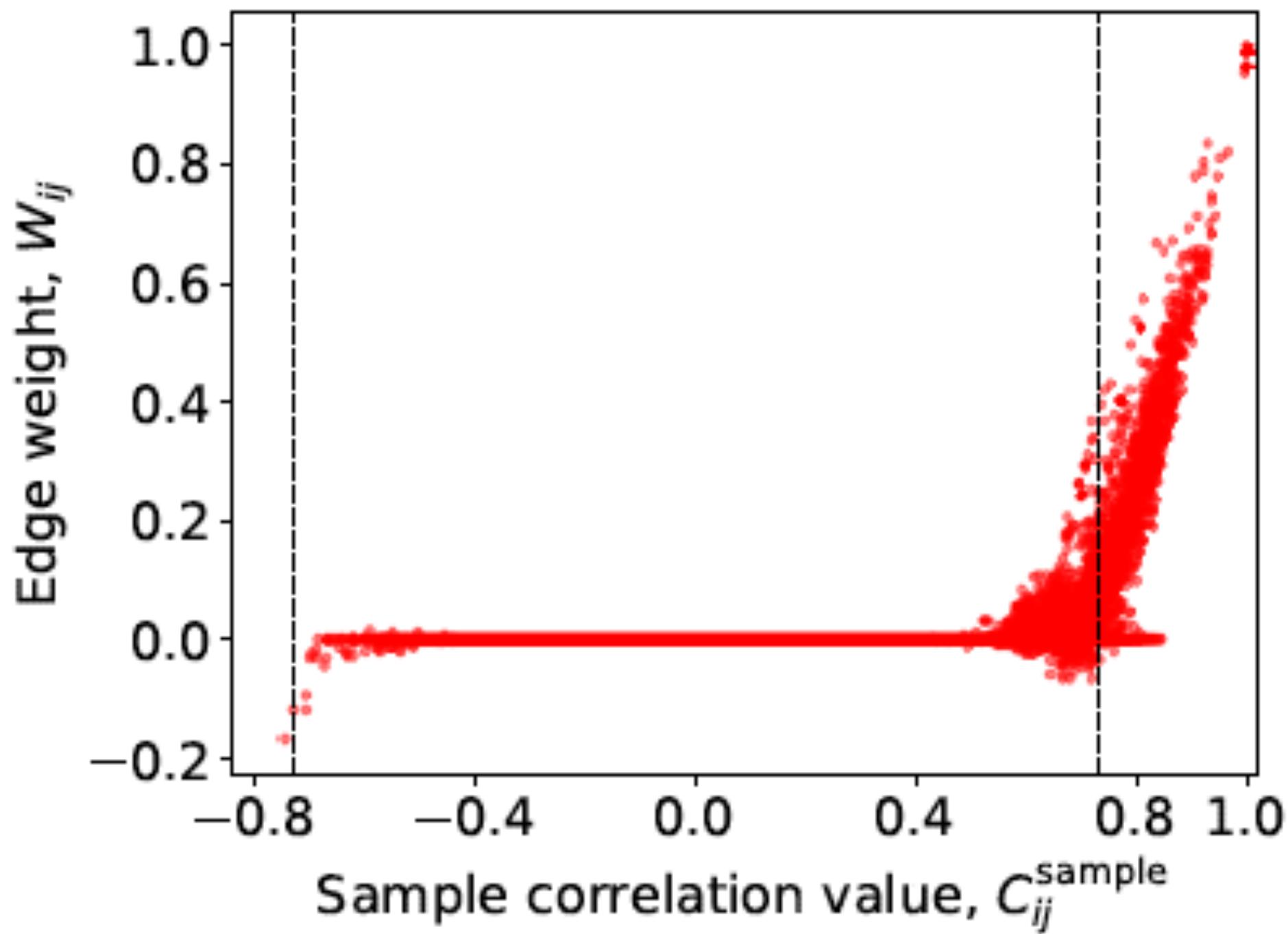
Adaptive Lasso:

$$\lambda_{ij} = \bar{\lambda} |W_{ij}^{\text{MLE}}|^{-\gamma}$$

where $\bar{\lambda} \geq 0$ and $\gamma > 0$ are hyperparameters.

Determined by a
model selection
criterion.

We use a typical
value $\gamma = 2$.



Model selection

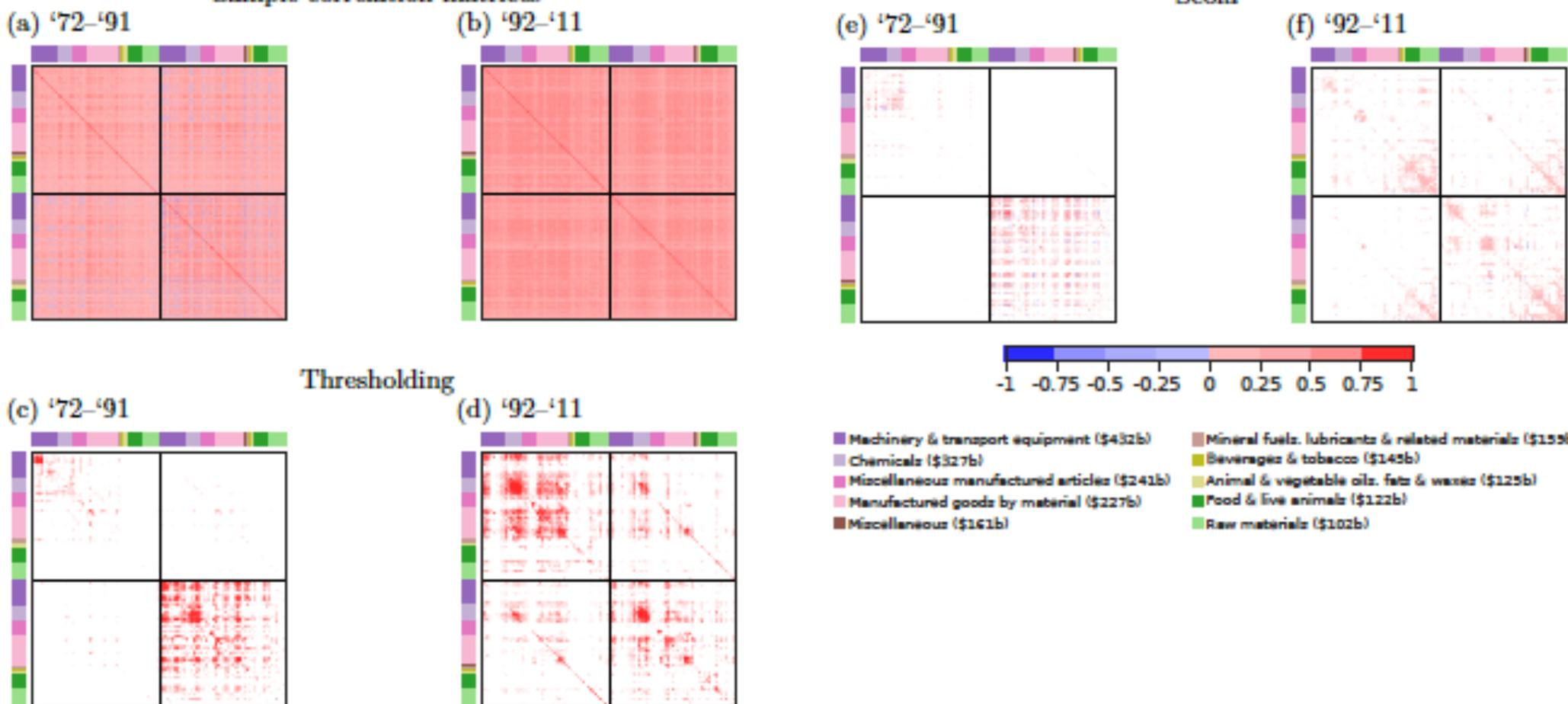
- Extended Bayes Information Criterion (Chen & Chen, Biometrika, 2008; Foygel & Drton, NIPS'10)
 - Smaller the better.
- WN: White noise null model
- HQS: Hirschberger-Qi-Steuer null model (\approx Erdős-Rényi RG)
- con: Our configuration model

Data	Null model					
	Correlation matrix			Precision matrix		
	C^{WN}	C^{HQS}	C^{con}	C^{WN}	C^{HQS}	C^{con}
Product space						
'72-'91	0.226	0.202	0.187	0.232	0.190	0.208
'92-'11	0.300	0.301	0.226	0.304	0.238	0.256
S&P 500						
'00-'07	0.600	0.562	0.553	0.640	0.557	0.587
'08-'15	0.664	0.558	0.524	0.596	0.539	0.607
Nikkei	0.991	0.874	0.861	1.001	0.837	0.882

Application to product space data/networks

- Prediction

Sample correlation matrices



- → Poster by Sadamori Kojaku

Conclusions

- Configuration model for correlation/covariance matrices
 - Maximum entropy
 - Node strength (i.e. weighted degree) preserved
 - Illustration with clustering coefficients and community detection
- Filtering as an alternative to thresholding
- Code/algorithmic documentation on Github (in python and MATLAB)
- Future problems: more examples, wherever null models are needed, ...

INTERNATIONAL SCHOOL AND CONFERENCE ON NETWORK SCIENCE

~Networks and Innovation~

January 20-23, 2020

International Conference Center

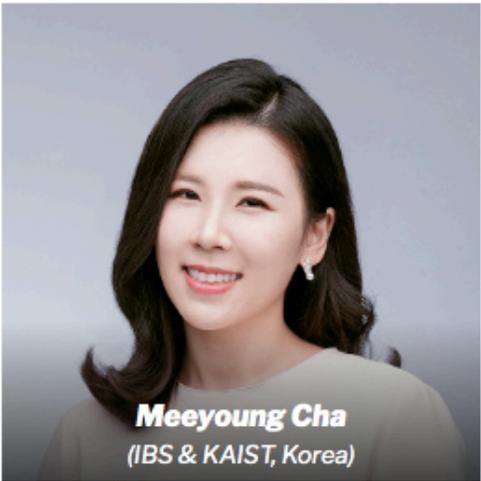
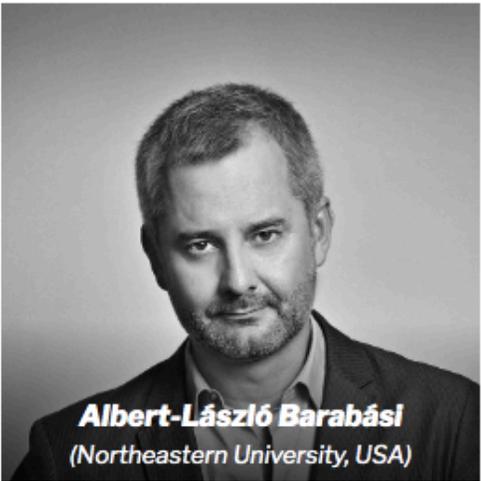
Waseda University

Tokyo, Japan

Hosted by Waseda Innovation Lab

Submission:
1 July - 30 September

Keynotes:



Invited:



Acknowledgments

Sadamori Kojaku
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postdoc at University of Bristol)



Yukie Sano
(Tsukuba University)



References

- Masuda, Kojaku & Sano, Physical Review E, 98, 012312 (2018).
- Kojaku & Masuda, arXiv:1903.10805